

COMPETING EFFECTS DESIGNS AND MODELS FOR TWO-DIMENSIONAL FIELD ARRANGEMENTS

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SUMMARY

Three methods of constructing *balanced* nearest neighbor row-column or competition effect designs are presented. Statistical response models for two-dimensional layouts and competing effects are formulated and corresponding statistical analyses are developed for designs of this type. The problem of obtaining solutions for all competing effects is discussed and illustrated. Two numerical examples illustrating aspects of design and statistical analyses are presented; one is an actual experiment and the other is artificial to demonstrate effect of competition on estimates of parameters. The problems of appropriate borders, spatial arrangements, measuring and/or eliminating competition effects, and the effect of not being able to obtain estimable contrasts among the competition effects are discussed.

1. Introduction

Various spatial statistics can take account of correlation between adjoining experimental units (plots) in field experiments. Nearest neighbor and competition effects designs and analyses for one-dimensional layouts were considered in papers edited by Kempton (1984). Also, Kempton (1982), Kempton and Lockwood (1984), and Besag and Kempton (1986), among others, considered statistical analyses for competition effects. Kempton (1982)

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proposed a single degree of freedom for competition effects which is a diagnostic statistic similar to Tukey's one-degree-of-freedom for nonadditivity. The present paper considers designs and statistical analyses for assessing the competition effect of individual treatments in two-dimensional layouts of field experiments. Three methods of constructing row-column designs for v treatments are given. The designs are balanced for competition effects. They are also balanced for nearest neighbor analysis but that is not the topic of this paper. Two numerical examples illustrate various aspects of design and statistical analysis.

2. Construction of Two-Dimensional Designs

Various row-column designs have been constructed which take account of the number and position of neighbors. Most were constructed for plant breeding experiments where, for the purposes of pollination, a cultivar is bordered by all its neighbors. A historical account and description of these may be found in Freeman (1979a, 1979b, 1981). He also showed similarities and differences between various designs and illustrated how to construct some nearest neighbor squares and rectangle designs. Further row-column designs constructed in the repeated measures context (a non-directional nearest neighbor in the Freeman sense) were given by Hedayat and Federer (1984).

Three methods for constructing experiment designs *balanced* for competition effects are given below. Latin square designs balanced for one-period carry-over or residual effects (see Williams, 1949, and Bradley, 1958, e.g.) may be used to construct latin square designs balanced for competition effects. Such latin squares have been denoted as row complete in Denes and Keedwell (1974). They denote a latin square design which is balanced for residual effects in *both* rows and columns as a complete latin square; in such squares, a treatment will occur equally often next to each of the other treatments but will not appear next to itself. The design is said to be balanced for competition effects. The idea of balance may be used in any two-dimensional layout and not solely for row-column designs.

Construction Method I. Utilizing the results in Sections 2.3 and 3.1 of Denes and Keedwell (1974), row-column designs may be constructed such that they are balanced for competition effects of adjacent plots. Complete latin squares may be constructed as follows. Let $v = 2m$, for m any positive integer, and rearrange the rows of a cyclic latin square of order v to have the first element of the rows as

a) $0, 1, 2m-1, 2, 2m-2, 3, 2m-3, \dots, m+1, m$, or as

b) $0, 2m-1, 1, 2m-2, 2, 2m-3, 3, \dots, m-1, m$.

A latin square in either the a) or b) arrangement is balanced for residual effects or is row complete. Rearranging the columns by either a) or b) results in a complete latin square balanced for competition effects of a treatment with the other treatments. Other methods for constructing complete latin squares may be found in Freeman (1979b). These designs are also nearest-neighbor designs and are denoted as $NND(v,v;0,4)$ where 0 is the number of times a treatment borders itself and 4 is the number of times a treatment is adjacent to each of the other treatments.

Construction Method II. Construct an F-square $FS(2v;2^v)$ (see Hedayat and Seiden, 1970) as the Kronecker product of a J matrix (all ones) of side 2 and a cyclic latin square of order v , any integer, i.e., $J_2 \otimes LS(v)$ where \otimes denotes Kronecker product. Then apply the procedure of Construction Method I to this F-square. The designs obtained by this method are denoted as $NND(2v,v;8,16)$ designs where $2v$ refers to the order of the F-square, v is the number of treatments, 8 is the number of times a treatment appears next to itself, and 16 is the number of times a given treatment is bordered by a treatment other than itself. Instead of J_2, J_3, J_4 , etc. could also be used to develop other F-squares.

For $v=3$ and using procedure b) of I results in

$$\begin{array}{ccccccc}
 & F(6;2,2,2) & & ABCABC & & ABCABC & & ACBBCA \\
 & & & BCABCA & & CAB CAB & & CBAABC \\
 & & & CAB CAB & & BCABCA & & BACCAB \\
 \left(\begin{array}{cc} 1 & 1 \\ 1 & 1 \end{array} \right) \otimes \left[\begin{array}{c} ABC \\ BCA \\ CAB \end{array} \right] & \rightarrow & & ABCABC & \rightarrow & BCABCA & \rightarrow & BACCAB = NND(6,3;8,16), \\
 & & & BCABCA & & CAB CAB & & CBAABC \\
 & & & CAB CAB & & ABCABC & & ACBBCA
 \end{array}$$

where \otimes denotes a symbolic Kronecker product. The last plan above is a nearest neighbor design for three treatments in an F-square of order 6. Treatment i , $i \neq i'$, appears next to treatment i' 16 times and treatment i borders itself 8 times.

Construction Method III. If a cultivar is to border itself for Construction Method I designs, simply repeat the last row of the design, as Patterson and Lucas (1959) did for repeated measures designs. Each of the v treatments would then appear $v+1$ times in the design. The treatments would remain orthogonal to rows but would be in a balanced block arrangement (see Shafiq and Federer, 1979) with respect to columns.

To illustrate Construction Method I and procedure b) on both rows and columns for $v=4$,

LS(4)		RM(4)		NN(4;4;0,4)
ABCD		ABCD		ADBC
BCDA	\rightarrow	DABC	\rightarrow	DCAB
CDAB		BCDA		BACD
DABC		CDAB		CBDA

An extra row on the above design with the last row repeated produces NNDs with $v+1$ rows and v columns for v treatments and every treatment is bordered by every other treatment including itself.

Any row-column design may be used as a nearest neighbor design. However, not all contrasts among a set of effects may be estimable. For most designs, the Kempton (1982) single degree of freedom for competition should be estimable even if solutions for all competition effects may not be obtainable. We describe some of the effects of not being able to estimate all competition effects in Example 2.

3. Response Model Equations for Block and Row-Column Designs

For a model taking into account competing effects of four neighbors in a row-column design, consider the following equation for the response from the plot in row f and column g whose treatment is h :

$$Y_{fghijkm} = \mu + \rho_f + \gamma_g + \tau_h + \alpha_i + \alpha_j + \alpha_k + \alpha_m + \epsilon_{fghijkm}, \quad (1)$$

where μ is a general mean effect, ρ_f is the effect of row f ($f=1, \dots, r$), γ_g is the effect of column

g ($g=1, \dots, c$), τ_h is the effect of treatment h ($h=1, \dots, v$), α_p is the competition effect of treatment $p=1, \dots, v$ in four adjacent positions ($i, j, k, m=1, \dots, v, x$), where x is an outside border for the row-column design, and $\epsilon_{fghijk m}$ are random error effects distributed with mean zero and common error variance σ_ϵ^2 . The competition effect of any treatment is independent of the position which it occupies. For a blocked design where each plot has four neighbors, simply delete ρ_f and f from the above equation and let γ_g denote the block effect. Also, if each plot has only two neighbors, delete α_k , α_m , k , and m from (1) to obtain the response model equation.

Diagonal neighbors are not considered because, for rectangular plots, they touch plot fgh on a corner only, and hence can cause little or no competition with it. Also, model (1) is most plausible when plots are square. When plots are rectangular with dimensions of length ℓ and width w , the weights $4\ell/(2\ell + 2w)$ might reasonably be given to the two α_p on the longer sides of the plot and weights $4w/(2\ell + 2w)$ on the shorter sides. If different spacings are used, replace ℓ by $\ell + d_1$ and w by $w + d_2$, where d_1 and d_2 are spacing widths between plots.

In row-column designs a solution for α_x , the competition effect of the border x , is not possible since α_x is completely confounded with the first and last rows and the first and last columns. Despite this it may be useful to leave α_x in (1) in order to understand which estimable contrasts are affected by an outside border effect. We omit α_x in our solutions and numerical examples. In order to make the effect of a border equal to zero, a composite of equal amounts of all the treatments in an experiment may perhaps be used as the border x . The solutions for competition effects add to zero and hence a border with all competition effects equally represented should not exhibit a competition effect. The use of a single cultivar x could exhibit a nonestimable competition effect on all border plots.

Solutions for row, column, treatment, and competition effects subject to the usual constraints, for the complete latin squares obtained from Construction Method I are possible when $v \geq 8$. From Construction Method II, solutions exist for $v \geq 4$. For $v=6$, none of the six complete squares given by Freeman (1979b) result in solutions for all effects. Solutions are

obtainable when $v = 4$ or 6 if only a row (column) complete latin square or one from Construction Method III is used; the design will not be balanced for competition effects.

Instead of explicit solutions for the row-columns designs obtained by the three construction methods, the following general solution is used because of the numerous designs involved and of the wide availability of PC computer software such as GAUSS, GENSTAT and SAS. This procedure handles all situations including unbalanced designs and missing observations. Using the standard linear model notation, for equation (1), let the design matrix be X . Subtracting the matrix

$$P_{a \times a} = \begin{bmatrix} 0_{a \times 1} & r1_{1 \times r} & c1_{1 \times c} & v1_{1 \times v} & b1_{1 \times v} \\ \hdashline & 0_{r \times r} & J_{r \times (c+2v)} & & \\ \hdashline & J_{c \times r} & 0_{c \times c} & J_{c \times 2v} & \\ \hdashline & J_{v \times (r+c)} & 0_{v \times v} & J_{v \times v} & \\ \hdashline & J_{v \times (r+c+v)} & & 0_{v \times v} & \\ \vdots & & & & \end{bmatrix}, \quad (2)$$

for $a = r + c + 2v + 1$ from $X'X$, we obtain $Z = X'X - P$ which will have an inverse when solutions for all effects are possible. For $b = r + c + v + 1$, let

$$Z_{a \times a} = \begin{bmatrix} D_{b \times b} & A_{b \times v} \\ N_{v \times b} & F_{v \times v} \end{bmatrix}, \quad (3)$$

$$X'Y = \begin{bmatrix} W_{1, b \times 1} \\ W_{2, v \times 1} \end{bmatrix}, \quad \text{and} \quad (4)$$

$$\hat{\beta} = Z^{-1}X'Y. \quad (5)$$

The last equation above produces solutions for the $1 + r + c + 2v$ effects when Z has an inverse. The solutions for $\hat{\alpha}$ may also be obtained as

$$\hat{\alpha} = (F - ND^{-1}A)^{-1}(W_2 - ND^{-1}W_1). \quad (6)$$

The sum of squares for the competition effects eliminating all else may be computed as

$$\hat{\alpha}'(\mathbf{W}_2 - \mathbf{N}\mathbf{D}^{-1}\mathbf{W}_1) . \quad (7)$$

Likewise, the treatment effects and sums of squares for treatments eliminating all other effects may be computed as follows. Rearrange \mathbf{Z} by interchanging the columns for treatment effects and the columns for competition effects. Then, interchange the rows for treatment effects and those for competition effects. The resulting matrix may be represented as

$$\mathbf{Z}^+ = \begin{bmatrix} \mathbf{E}_{b \times b} & \mathbf{G}_{b \times v} \\ \mathbf{H}_{v \times b} & \mathbf{K}_{v \times v} \end{bmatrix} ;$$

also partition the $\mathbf{X}'\mathbf{Y}$ vector as $\begin{bmatrix} \mathbf{W}_3 \\ \mathbf{W}_4 \end{bmatrix}$. Then a solution for treatment effects is

$$\hat{\tau} = (\mathbf{K} - \mathbf{H}\mathbf{E}^{-1}\mathbf{G})^{-1}(\mathbf{W}_4 - \mathbf{H}\mathbf{E}^{-1}\mathbf{W}_3) . \quad (8)$$

The corresponding sum of squares is computed as

$$\hat{\tau}'(\mathbf{W}_4 - \mathbf{H}\mathbf{E}^{-1}\mathbf{W}_3) . \quad (9)$$

4. Illustrative Examples

We now present examples illustrating the computational procedure. Example 1 demonstrates estimation of competition effects for a latin square of order 6, with allowance for a rectangular experimental unit. Here the treatment sum of squares eliminating all else is much reduced from the treatment ignoring competition effects sum of squares. The variances for contrasts are described. Example 2 illustrates the method of analysis for designs constructed by Method II. The data are artificial and residuals have been included. Thus, the sum of squares for the “error” line in the ANOVA is known, i.e., the sum of the squared known residuals. This design by Construction Method II for $v=3$ does not allow solutions for all competition effects. The biasing effect of not being able to estimate all competition effects is illustrated.

Example 1 Das and Giri (1979), page 77, gave the following field layout and data for a wheat experiment in a latin square design of order six:

Column (yields in grams/10)							
Row	1	2	3	4	5	6	Total
1	219 F	250 E	227 D	162 C	182 B	91 A	1131
2	227 E	141 C	91 A	191 D	213 F	195 B	1058
3	204 B	91 A	225 F	229 E	250 D	207 C	1206
4	77 A	204 B	240 E	199 F	182 C	250 D	1152
5	250 D	231 F	209 C	204 B	91 A	227 E	1212
6	152 C	186 D	191 B	77 A	230 E	198 F	1034
Total	1129	1103	1183	1062	1148	1168	6793

The treatment totals are: $Y_{..A.} = 518$, $Y_{..B.} = 1180$, $Y_{..C.} = 1053$, $Y_{..D.} = 1354$, $Y_{..E.} = 1403$, $Y_{..F.} = 1285$. The competition effect totals are: $Y_{...A} = 4006.82$, $Y_{...B} = 2899.90$, $Y_{...C} = 4087.28$, $Y_{...D} = 3796.82$, $Y_{...E} = 3675.44$, $Y_{...F} = 4130.22$. The solutions for the effects are:

$$\begin{aligned}
 \hat{\mu} &= 188.694 \\
 \hat{\rho}_1 &= -1.063 & \hat{\gamma}_1 &= -0.452 & \hat{\tau}_A &= -102.262 & \hat{\alpha}_A &= -5.947 \\
 \hat{\rho}_2 &= -12.478 & \hat{\gamma}_2 &= -4.750 & \hat{\tau}_B &= 14.985 & \hat{\alpha}_B &= -2.039 \\
 \hat{\rho}_3 &= 11.696 & \hat{\gamma}_3 &= 8.497 & \hat{\tau}_C &= -16.397 & \hat{\alpha}_C &= 0.074 \\
 \hat{\rho}_4 &= 2.425 & \hat{\gamma}_4 &= -11.831 & \hat{\tau}_D &= 36.423 & \hat{\alpha}_D &= 3.106 \\
 \hat{\rho}_5 &= 14.784 & \hat{\gamma}_5 &= 2.630 & \hat{\tau}_E &= 43.355 & \hat{\alpha}_E &= 1.662 \\
 \hat{\rho}_6 &= -15.363 & \hat{\gamma}_6 &= 5.907 & \hat{\tau}_F &= 23.898 & \hat{\alpha}_F &= 3.145
 \end{aligned}$$

The treatments are: A = no nitrogen (N), B = 40 kg. N/hectare, C = 80 kg./hectare, D = 120 kg./hectare, E = 160 kg./hectare, F = 200 kg./hectare. The plot size was 8×0.6 meters.

The total perimeter of a plot was $2(8+0.6) = 17.2$ meters. Therefore the coefficients for competition effects in the design matrix are $4(8/17.2) = 1.86$ and $4(0.6/17.2) = 0.14$. That is, the sides of the experimental unit are bordered by a neighbor for a length of 8 meters, whereas the ends are bordered by only 0.6 meters. The design matrix is $X = [X_1 \ X_2]$ where X_1 is the usual design matrix for a latin square of order six and

$$X_{2,25 \times 6} = \begin{bmatrix} 0 & 0 & 0 & 0 & 2 & 0 \\ 0 & 0 & .14 & 1.86 & 0 & 1.86 \\ .14 & 0 & 1.86 & 0 & 1.86 & 0 \\ 0 & 1.86 & 0 & 2 & 0 & 0 \\ 1.86 & 0 & 1.86 & 0 & 0 & .14 \\ 0 & 2 & 0 & 0 & 0 & 0 \\ & & & \vdots & & \\ 0 & 0 & 0 & 2 & 0 & 0 \\ 0 & 1.86 & 1.86 & 0 & 0 & .14 \\ 1.86 & 0 & .14 & 1.86 & 0 & 0 \\ 0 & 2 & 0 & 0 & 1.86 & 0 \\ 2 & 0 & 0 & 0 & 0 & 1.86 \\ 0 & 0 & 0 & 0 & 2 & 0 \end{bmatrix}.$$

An ANOVA for this example is:

Source of variation	df	Sum of squares	Mean square
Total	38	1,381,069	—
Correction for mean	1	1,281,801.36	—
Row (ignoring competition)	5	4,559.47	911.89
Column (ignoring competition)	5	1,650.47	330.09
Nitrogen levels (ignoring competition)	5	88,612.47	—
Remainder	20	4,445.22	222.26
Competition (eliminating all else)	5	1,102.68	220.54
Error	15	3,342.54	222.84
Nitrogen (eliminating all else)	5	53,114.90	10,622.98

The sum of squares for competition effects eliminating all other effects is computed as:

$$\hat{\alpha}'(W_2 - ND^{-1}W_1) = \begin{bmatrix} -5.949 \\ -2.039 \\ 0.074 \\ 3.106 \\ 1.663 \\ 3.145 \end{bmatrix}' \cdot \begin{bmatrix} -108.946 \\ -19.792 \\ 11.978 \\ 89.478 \\ -33.456 \\ 60.738 \end{bmatrix} = 1,102.68.$$

The sum of squares for treatments eliminating all other effects is computed as:

$$\hat{\tau}'(W_4 - HE^{-1}W_3) = \begin{bmatrix} -102.262 \\ 14.986 \\ -16.398 \\ 36.422 \\ 43.355 \\ 23.896 \end{bmatrix}' \cdot \begin{bmatrix} -373.601 \\ -41.860 \\ -29.990 \\ 111.456 \\ 154.406 \\ 179.590 \end{bmatrix} = 53,114.90.$$

The estimated variance-covariance matrix for $\hat{\alpha}$ is¹

$$\hat{\sigma}_\epsilon^2(\mathbf{F} - \mathbf{ND}^{-1}\mathbf{A})^{-1} = 222.84 \begin{bmatrix} 0.058 & -0.001 & 0.001 & -0.002 & -0.023 & -0.004 \\ -0.001 & 0.053 & -0.006 & 0.000 & -0.000 & -0.017 \\ 0.001 & -0.006 & 0.054 & -0.020 & -0.013 & 0.013 \\ -0.002 & 0.000 & -0.020 & 0.050 & 0.015 & -0.014 \\ -0.023 & -0.000 & -0.013 & 0.015 & 0.054 & -0.003 \\ -0.004 & -0.017 & 0.013 & -0.014 & -0.003 & 0.055 \end{bmatrix}.$$

The estimated variance-covariance matrix for $\hat{\tau}$ is¹

$$\hat{\sigma}_\epsilon^2(\mathbf{K} - \mathbf{HE}^{-1}\mathbf{G})^{-1} = 222.84 \begin{bmatrix} 0.275 & -0.065 & 0.027 & -0.015 & -0.037 & 0.034 \\ -0.065 & 0.283 & -0.023 & 0.019 & 0.040 & -0.036 \\ 0.027 & -0.023 & 0.235 & -0.052 & 0.001 & 0.029 \\ -0.015 & 0.019 & -0.052 & 0.251 & 0.023 & -0.008 \\ -0.037 & 0.040 & 0.001 & 0.023 & 0.231 & -0.039 \\ 0.034 & -0.036 & 0.029 & -0.008 & -0.039 & 0.239 \end{bmatrix}.$$

Something appears to be amiss with the data for this example. Only two values, 77 and 91, were obtained for treatment A, 9 of the 36 values are integral multiples of one pound, 15 of the 36 are multiples of one-half pound, and if the remaining ones should be multiples of one-fourth pound, some arithmetic errors were made in converting to grams. Also, it appears that the residual sum of squares may be too large, perhaps because of nonadditivity, heterogeneous variances for treatments, or because gradients run diagonally through the square. However, one fact is clear from the data: either of the contrasts A versus the rest or a linear trend of effects on nitrogen level would account for a large proportion of the treatment (eliminating all other effects) and competition (eliminating all other effects) sums of squares. These sums of squares, respectively, are 45,846.17 and 777.74 for the first contrast. These were computed as $\ell\hat{\alpha}\ell(\mathbf{W}_2 - \mathbf{ND}^{-1}\mathbf{W}_1)/\ell\ell'$; where $\ell = (-5 \ 1 \ 1 \ 1 \ 1 \ 1)$. Even though the "Error" sums of squares is probably too large, the single degree of freedom sum of squares for competition effects is significant at about the 8% level. Biologically, this

¹ Various computer programs give different variance-covariance matrices. However, adding an appropriate constant (of the order of .01) times a J matrix (all ones) to any one of them yields another. The variance of a contrasts is identical for all forms obtained. We obtained three different forms.

would be explainable by the fact that when nitrogen was available, plants got off to an earlier start and were larger. The larger plants made use of the nutrients in the adjoining plots with smaller plants. A linear trend, i.e., $\ell = (-5 \ -3 \ -1 \ 1 \ 3 \ 5)$, illustrated somewhat the same effect on the sums of squares, i.e., they were 38,394.11 and 753.55. The remainder mean square for nitrogen effects after removal of the contrast A versus rest sum of squares was 1817.18 with four degrees of freedom, resulting in an F-value of 8.2. A linear regression among the nonzero levels of nitrogen would account for a large proportion of this variation.

Some items to note are:

- i) The sum of squares for treatment eliminating all other effects is only $3/5$ ths as large as the treatment ignoring competition effects sum of squares.
- ii) Plot shape needs to be considered in assessing competition effects.
- iii) A linear regression of treatment and competition effects on amount of nitrogen applied would account for a considerable and significant proportion of their respective sums of squares.
- iv) The inferences for fertilizer effect would be little affected with regard to slope of fertilizer treatment responses to increasing nitrogen even if competition were ignored.
- v) The design matrix X has coefficients other than 0 and 1 and some are not integers.
- vi) The variances for competition contrasts are much smaller than for fertilizer treatment contrasts. The extra replication for competition effects accounts for this.

Example 2 An artificial example is constructed from the following values for the effects for a design from Construction Method II for $v=3$ treatments and response model equation (1):

$$\begin{array}{llll}
 \mu = 20 & & & \\
 \rho_1 = -5 & \gamma_1 = -3 & \tau_A = -7 & \epsilon_{12C} = -1 \\
 \rho_1 = 5 & \gamma_2 = -3 & \tau_B = 2 & \epsilon_{13B} = 1 \\
 \rho_3 = 0 & \gamma_3 = 6 & \tau_C = 5 & \epsilon_{21C} = 1 \\
 \rho_4 = 0 & \gamma_4 = 0 & \alpha_A = -3 & \epsilon_{23A} = -1 \\
 \rho_5 = 0 & \gamma_5 = 0 & \alpha_B = -1 & \epsilon_{31B} = -1 \\
 \rho_6 = 0 & \gamma_6 = 0 & \alpha_C = 4 & \epsilon_{32A} = 1
 \end{array}$$

All other ϵ_{fghijk} , or residuals, are set equal to zero. The sum of residuals is zero over rows, columns, treatments, and competition effects. From the above values the data and design are:

Rows	Columns (yields and treatments)						Total
	1	2	3	4	5	6	
1	13A	11C	24B	17B	15C	16A	96
2	23C	26B	22A	17A	29B	25C	142
3	18B	10A	33C	27C	12A	22B	122
4	19B	9A	33C	27C	12A	22B	122
5	17C	21B	18A	12A	24B	20C	112
6	18A	17C	28B	22B	20C	21A	126
Total	108	94	158	122	112	126	720

For this design not all effects have solutions with the constraints $\Sigma \hat{\rho}_f = \Sigma \hat{\gamma}_g = \Sigma \hat{\tau}_h = \Sigma \hat{\alpha}_i = 0$. To determine which linear contrasts of the $\hat{\alpha}$ are not estimable, one may obtain the eigenvalues and eigenvectors for the matrix $(F - ND^{-1}A)$. Or, one can go back to the original $X'X$ matrix, and apply the constraints for the row, column, and treatment effects but not for the competition effects. Let the matrix $X'X$ with the above constraints be

$$Z_{a \times a}^* = \begin{bmatrix} D_{b \times b} & A_{b \times v}^* \\ N_{v \times v}^* & F_{v \times v} \end{bmatrix}, \quad (10)$$

where D and F have been defined previously and A^* and N^* are the corresponding submatrices of $X'X$. In this form the eigenvalues and eigenvectors of $(F - N^*D^{-1}A^*)$ can be obtained and examined. Omitting the last constraint but using the first three on $X'X$, denoting the resulting matrix as Z^* as in equation (10), the following matrix is obtained:

$$F - N^*D^{-1}A^* = \begin{bmatrix} 0.8889 & -4.4444 & 3.5556 \\ -4.4444 & 22.2222 & -17.7778 \\ 3.5556 & -17.7778 & 14.2222 \end{bmatrix},$$

where $Z^* = \begin{bmatrix} D & A^* \\ N^* & F \end{bmatrix}$. The nonzero eigenvalue of the above matrix is $37\frac{1}{3}$ and the corresponding eigenvector is $(0.15430 \ -0.77152 \ 0.61721) = E'$. Since only one eigenvalue is nonzero this indicates that the total sum of squares for competition effects is obtained from the single degree of freedom contrast $0.15430\alpha_A - 0.77152\alpha_B + 0.61721\alpha_C = Q_1$. Now, form the matrix $Z_1 = \begin{bmatrix} D & A^*E \\ E'N^* & E'FE \end{bmatrix}$ where $D_{16 \times 16}$ has the following vector as its first column and main right diagonal:

$$(36 \ 6 \ 6 \ 6 \ 6 \ 6 \ 6 \ 6 \ 6 \ 6 \ 6 \ 6 \ 6 \ 6 \ 12 \ 12 \ 12) .$$

$$(A^*E)'_{1 \times 16} = (E'N^*)_{1 \times 16} =$$

$$\begin{bmatrix} -0.0004 & -0.3087 & -1.2345 & 1.5430 & 1.5430 & -1.2345 & -0.3087 & -0.3087 & -1.2345 & 1.5430 \\ 1.5430 & -1.2345 & -0.3087 & -1.2346 & 6.1720 & -4.9378 & & & & \end{bmatrix}$$

and $(E'FE)_{1 \times 1} = 45.3333$. Z_1 now has an inverse. Form totals $T = X'X$ where X is the incidence matrix and Y is the observation vector. Let T_1 be the first 16 totals from T plus E' times the last three totals of T . Then, the solutions for effects are obtained as $\hat{\beta}_1 = (Z_1)^{-1}T_1$ and are:

Solution	Bias	Solution	Bias
$\hat{\mu} = 20.000$	0	$\hat{\gamma}_3 = 5.619$	8/21
$\hat{\rho}_1 = -3.857$	-3(8/21)	$\hat{\gamma}_4 = -0.381$	8/21
$\hat{\rho}_2 = 4.238$	2(8/21)	$\hat{\gamma}_5 = -0.762$	2(8/21)
$\hat{\rho}_3 = -0.381$	8/21	$\hat{\gamma}_6 = 1.143$	-3(8/21)
$\hat{\rho}_4 = -0.381$	8/21	$\hat{\tau}_A = -4.714$	-6(8/21)
$\hat{\rho}_5 = -0.762$	2(8/21)	$\hat{\tau}_B = 1.238$	2(8/21)
$\hat{\rho}_6 = 1.143$	-3(8/21)	$\hat{\tau}_C = 3.476$	4(8/21)
$\hat{\gamma}_1 = -1.857$	-3(8/21)	$\hat{\alpha}_1 = 2.777$	0
$\hat{\gamma}_2 = -3.762$	2(8/21)		

Adding the bias to each of the solutions results in the effect values used to construct the example. The bias is a multiple of 8/21. Because this particular design did not allow solutions for $\hat{\alpha}_A$, $\hat{\alpha}_B$, and $\hat{\alpha}_C$, the other parameter estimates are biased.

To obtain the sum of squares for competition effects eliminating all other effects, proceed as before except that Z_1 and T_1 are used in place of Z and $X'Y$ and it has one degree of freedom.

The following ANOVA table was obtained using these solutions and the methods described above and in previous examples:

Source of variation	df	Sum of squares
Total	36	15,734
Correction for mean	1	14,400
Row (ignoring competition)	5	194.67
Column (ignoring competition)	5	394.67
Treatment (ignoring competition)	2	450.66
Remainder (ignoring competition)	23	294.00
Competition (eliminating all else)	1	288.00
Residual	22	6.00
Treatment (eliminating all else)	2	429.77

Some items to note are:

- i) This design does not allow solutions for all effects.
- ii) The solutions obtained for the effects are biased.
- iii) The solutions obtained ignoring competition effects would also have been biased. They are:

$$\begin{array}{lll}
 \hat{\mu} = 20 & & \\
 \hat{\rho}_1 = -4 & \hat{\gamma}_1 = -2 & \hat{\tau}_A = -5 \\
 \hat{\rho}_2 = 11/3 & \hat{\gamma}_2 = -13/3 & \hat{\tau}_B = 8/3 \\
 \hat{\rho}_3 = 1/3 & \hat{\gamma}_3 = 19/3 & \hat{\tau}_C = 7/3 \\
 \hat{\rho}_4 = 1/3 & \hat{\gamma}_4 = 1/3 & \\
 \hat{\rho}_5 = -4/3 & \hat{\gamma}_5 = -4/3 & \\
 \hat{\rho}_6 = 1 & \hat{\gamma}_6 = 1 &
 \end{array}$$

- iv) The estimated error variance for an F-square would have been $294/23$ and would be $6/22$ in the above ANOVA whereas it should have been $6/21$ as the example was constructed.
- v) The above illustrates that if competition effects are present, they should all be estimable as they were in Example 1. Otherwise, the results may be vitiated in that the parameter estimates are biased.

- vi) All F-tests are biased.
- vii) If a different arrangement of columns had been used, solutions for all competition effects may have been possible.

5. Discussion

Competition can be important in field experiments. The first author has seen experiments in maize wherein one cultivar was completely eliminated by its neighbors and in sugarcane where a neighboring cultivar had a very visible effect on the first two rows of a cultivar even though the rows were ten feet apart. Several other examples have been seen. Unless an experimenter desires to measure competitive effects (see, e.g., Jensen and Federer, 1964, 1965), he or she would be wise to use an arrangement of plots which would eliminate competition. This can be done through spacing or through using border rows. The latter utilizes additional space and material which are often limiting. Also, additional spacing may increase the heterogeneity within blocks, resulting in larger error mean squares. As repeatedly advocated by the late Dr. LeRoy Powers, geneticist and plant breeder, the use of space can eliminate the adverse effects of competition from adjoining plots and missing plants within a plot. In genetic studies, competitive effects must be eliminated from genetic effects whereas in commercial field arrangements intra-cultivar competitive effects must be considered in evaluating cultivars for sole cropping conditions. Likewise, inter-cultivar competitive effects must be considered for intercropping mixtures of cultivars. Experimenters have confused the goal of experiments and still do so. The *conditions of an experiment must emulate conditions to be used in practice* in order to make meaningful inferences. In sole cropping practices, inter-cultivar competition is not a factor and if it is present in an experiment evaluating cultivars, the inferences may be meaningless if competition is ignored.

In order to effectively eliminate inter-plot competition, we suggest that the rows between plots be approximately twice the distance of lateral root growth. For cereals, this is approximately the height of the plants. In order to obtain the same density per hectare, Dr.

Powers recommended that the density within plots be increased to satisfy this requirement. Another spatial arrangement for maize cultivars would be to have a plot of two rows, 0.25 meters apart and with 1.75 meters between pairs of rows. This arrangement would be comparable in density to one where rows are one meter apart and would effectively eliminate competition between plots for most maize cultivars. (Lambert (1983), e.g., used the paired row arrangement of maize in several experiments on sweet corn.) Other spatial arrangements are given in Federer (1990, Chapter 9) for sole and intercropping experiments.

With respect to border rows for an experiment, we suggest that a composite of equal amounts of all treatments be used. If the competition effects sum to zero, this composite would exert zero effect on the plots on the outsides of an experiment. However, competition effects need not sum to zero (see Federer, 1990, Chapters 6 and 7). As competition effects are estimated in the statistical analyses presented, they do sum to zero. Hence, a composite may be the answer for borders for some experiments, whether a varietal trial, a fertilizer, spraying, or other type of experiment. If the treatments are dates of planting, for example, an average date of planting could be used for the borders or the outside plots could be divided into equal areas with each area having one of the treatments.

The designs obtained by construction methods I, II, and III may not only be used for plant pollination studies as mentioned earlier, but are useful in plant association and plant competition studies. A plan obtained by Construction Method II has been used to study association and competition among five species in Australia.

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